

Financial time series

Gerencsér László

Vágó Zsuzsanna

Gerencsér Balázs

Solutions to Exercises

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Chapter 1

Basic concepts

1.1 Wide sense stationary processes

Exercise 1.1 Let (y_n) be a wide sense stationary process and let us define

$$u_n = a_1 y_{n-1} + \cdots + a_p y_{n-p}, \quad \text{with } a_k \text{ real, } k = 1, \dots, p.$$

Show that (u_n) is also a wide sense stationary process.

Solution: The mean value:

$$E(u_n) = E\left(\sum_{k=1}^p a_k y_{n-k}\right) = \sum_{k=1}^p a_k E(y_{n-k}) = 0.$$

The covariance of u_n and $u_{n+\tau}$ is the following:

$$\begin{aligned} Eu_{n+\tau}u_n &= E\left(\sum_{k=1}^p a_k y_{n-k+\tau} \sum_{l=1}^p a_l y_{n-l}\right) = \\ &= \sum_{k=1}^p \sum_{l=1}^p a_k a_l E(y_{n-k+\tau} y_{n-l}) = \sum_{k=1}^p \sum_{l=1}^p a_k a_l r(l - k + \tau). \end{aligned}$$

It is indeed independent of n . □

Exercise 1.2 Define the $p \times p$ matrix $R = (R_{k,l})$ by

$$R_{k,l} = r(l - k), \quad k, l = 1, \dots, p.$$

Set

$$Y = (y_{n-1}, \dots, y_{n-p})^T.$$

Prove that the matrix R defined above can also be written as $R = E(YY^T)$.

Solution: We have

$$YY^T = \begin{pmatrix} y_{n-1} \\ \vdots \\ y_{n-p} \end{pmatrix} \cdot (y_{n-1} \dots y_{n-p})^T,$$

thus $(YY^T)_{k,l} = y_{n-k}y_{n-l}$. Then

$$\mathbb{E}(YY^T)_{k,l} = \mathbb{E}y_{n-k}y_{n-l} = r(n-k-n+l) = r(l-k).$$

□

Exercise 1.3 Using the representation $R = \mathbb{E}(YY^T)$ prove that R is symmetric and positive semidefinite.

Solution: As the autocovariance function is symmetric, $(r(\tau) = r(-\tau))$, the matrix R is symmetric. We also have to prove, that for all $a \in \mathbb{R}^p$, the quadratic form $a^T R a \geq 0$. Indeed,

$$a^T R a = a^T \mathbb{E}(YY^T) a = \mathbb{E}a^T (YY^T) a = \mathbb{E}(a^T Y) (a^T Y)^T \geq 0.$$

□

Exercise 1.4 Let $y = (y_n)$ be a complex-valued w.s.st. process with autocovariance function $r^y(\cdot)$. Show that the matrix $R = (R_{k,l})$ defined by

$$R_{k,l} = r^y(l-k), \quad k, l = 1, \dots, p$$

is a Hermitian, positive semi-definite Toeplitz matrix.

Solution: As the entry $R_{l,k}$ only depends on $l-k$, it is a Toeplitz matrix. For complex w.s.st. processes we have

$$R_{k,l} = r^y(l-k) = \overline{r^y(k-l)} = \overline{R_{l,k}},$$

thus R is Hermitian. For all $a \in \mathbb{C}^p$ we have

$$\bar{a}^T R a = \bar{a}^T \mathbb{E}(\bar{Y}Y^T) a = \mathbb{E}\bar{a}^T (\bar{Y}Y^T) a = \mathbb{E}(\bar{a}^T \bar{Y}) (a^T Y)^T \geq 0.$$

□

1.2 Orthogonal processes and their transformations

Exercise 1.5 Show that the sequence $(y_{n,N})$ is a Cauchy sequence in L_2 norm.

Solution: As $y_{n,N} = \sum_{k=0}^N h_k e_{n-k}$, we have by the orthogonality of (e_k) for $N < M$

$$\|y_{n,M} - y_{n,N}\|^2 = \sum_{k=N+1}^M h_k^2 < \varepsilon,$$

if N is large enough, since by assumption $(h_k) \in \ell_2$. □

Exercise 1.6 Define y_n by

$$y_n = \sum_{k=0}^{\infty} h_k e_{n-k}.$$

Show that $y = (y_n)$ is a wide sense stationary process.

Solution: We know from the previous exercise, that

$$y_n = \lim_{N \rightarrow \infty} y_{n,N}.$$

Then

$$\mathbb{E}y_n = \lim_{N \rightarrow \infty} \mathbb{E}y_{n,N} = 0.$$

Also, for the autocovariance function we have

$$\lim_{N \rightarrow \infty, M \rightarrow \infty} \mathbb{E}y_{n+\tau,N}y_{n,M} = \sum_{k=0}^{\infty} h_k h_{k+\tau} \cdot \sigma^2.$$

□

1.3 Prediction

Exercise 1.7 Show that if R is singular then y_n can be predicted with 0 error.

Solution: If R is singular, then there is a nonzero vector $v = (v_1, \dots, v_p)^T$, such that $v^T R = 0$, and also $v^T R v = 0$. We can assume, that the first non-zero element of v equals to 1. Also, we can assume that $v = (1, v_2, \dots, v_p)^T$. Define

$$\hat{y}_n = \sum_{k=1}^p (-v_{k+1}) y_{n-k}.$$

Then

$$\begin{aligned} E(y_n - \hat{y}_n)^2 &= E \left(y_n + \sum_{k=1}^{p-1} v_{k+1} y_{n-k} \right)^2 = \\ &= v^T E(y_n, y_{n-1}, \dots, y_{n-p+1})^T (y_n, y_{n-1}, \dots, y_{n-p+1}) v = v^T R v = 0, \end{aligned}$$

thus the prediction error is 0. \square

Exercise 1.8 * Show that for the process (y_n^r) we have

$$H_{-\infty}^{y^r} = \{0\}.$$

Solution: Let us recall that

$$y_n^r = \sum_{k=0}^{\infty} c_k e_{n-k}, \quad c_0 = 1,$$

where $e = (e_n)$ is the innovation process of $y = (y_n)$. Therefore $y_n^r \in H_n^e$ for all n implying

$$H_n^{y^r} \subset H_n^e \quad \text{for all } n.$$

It follows that

$$H_{-\infty}^{y^r} \subset H_{-\infty}^e.$$

Now if $v \in H_{-\infty}^e$ then $v \in H_{-n}^e$ for all n . Since $e = (e_n)$ is an orthogonal process we have

$$e_m \perp H_{-n}^e \text{ for any } m > -n.$$

In particular

$$e_m \perp v.$$

Note that the latter claim holds independently of n , for all n . It follows that

$$v \perp H^e = H_{\infty}^e.$$

But, obviously, $v \in H_{\infty}^e$. Thus we must have $v = 0$.

Exercise 1.9 * Show that

$$H_{-\infty}^y = H_{-\infty}^{y^s}.$$

Solution: We have

$$y_n = y_n^r + y_n^s, \quad \text{for all } n.$$

The orthogonality $y^r \perp y^s$ implies

$$H_\infty^{y^r} \perp H_\infty^{y^s},$$

and thus

$$H_{-n}^{y^r} \perp H_{-n}^{y^s}, \quad \text{for all } n.$$

First we show that

$$H_{-\infty}^y \supset H_{-\infty}^{y^s}.$$

To see this note that the definition

$$e_n = y_n - (y_n | H_{n-1}^y)$$

implies $e_n \in H_n^y$ for all n , which in turn implies

$$H_n^e \subset H_n^y \quad \text{for all } n.$$

But then $y_n^r \in H_n^y$ implies

$$y_n^r \in H_n^y, \quad \text{for all } n,$$

and hence $y_n^s = y_n - y_n^r$ we also have

$$y_n^s \in H_n^y, \quad \text{for all } n.$$

We conclude from here that

$$H_n^{y^s} \subset H_n^y \quad \text{for all } n,$$

which implies

$$H_{-\infty}^{y^s} \subset H_{-\infty}^y.$$

Now to prove the opposite inclusion let $v \in H_{-\infty}^y$. Then $v \in H_{-n}^y$ for all n . The decomposition of y as $y_n = y_n^r + y_n^s$ with $y^r \perp y^s$ implies

$$H_{-n}^y = H_{-n}^{y^r} \oplus H_{-n}^{y^s}$$

for all n , where \oplus denotes orthogonal direct sum. Let us write v as

$$v = v_{-n}^r + v_{-n}^s$$

with $v_{-n}^r \in H_{-n}^{y^r}$ and $v_{-n}^s \in H_{-n}^{y^s}$. Then

$$v_{-n}^r = (v | H_{-n}^{y^r}).$$

Now letting n tend to infinity we get that

$$\lim_{n \rightarrow \infty} v_{-n}^r = \lim_{n \rightarrow \infty} (v|H_{-n}^{y^r}) = (v|H_{-\infty}^{y^r})$$

But $H_{-\infty}^{y^r} = \{0\}$, hence $v_{-n}^r \rightarrow 0$ and we conclude that

$$v = \lim_{n \rightarrow \infty} v_{-n}^s.$$

Obviously, the right hand side belongs to $H_{-m}^{y^s}$ for all m and hence it belongs to $H_{-\infty}^{y^s}$. \square

Chapter 2

Prediction, innovation and the Wold decomposition

2.1 Prediction using the infinite past

Exercise 2.1 *Prove that*

$$(y_n | H_{n-1}^y) = \lim_{p \rightarrow \infty} (y_n | H_{n-1, n-p}^y) \quad \text{in } L_2(\Omega, \mathcal{F}, P).$$

Solution: Let us denote

$$\hat{y}_n = (y_n | H_{n-1}^y), \quad \text{and} \quad \tilde{y}_n = y_n - \hat{y}_n.$$

Thus

$$y_n = \tilde{y}_n + \hat{y}_n,$$

where $\tilde{y}_n \perp H_{n-1}^y$, and also $\tilde{y}_n \perp (y_n | H_{n-1, n-p}^y)$. Thus projecting both sides onto $H_{n-1, n-p}^y$ we get

$$(y_n | H_{n-1, n-p}^y) = (\hat{y}_n | H_{n-1, n-p}^y).$$

But $\hat{y} \in \text{cl}(\cup_p H_{n-1, n-p}^y)$ implies

$$\lim_{p \rightarrow \infty} ((\hat{y}_n | H_{n-1, n-p}^y) - \hat{y}) = 0.$$

□

2.2 Singular processes

Exercise 2.2 *Show that if $H_n^y = H_{n-1}^y$ for a single n , then $H_n^y = H_{n-1}^y$ for all n .*

Solution: We know, that the process (e_n) with $e_n = y_n - (y_n | H_{n-1}^y)$ is a w.s.st. orthogonal process. If $H_n^y = H_{n-1}^y$, then one of these random variable is zero, thus all of them are zero. \square

Exercise 2.3

$$y_n = \xi e^{in\omega}, \quad n = 0, \pm 1, \pm 2, \dots,$$

Show that the above process is wide sense stationary.

Solution:

$$E y_n = E \xi e^{in\omega} = 0,$$

and

$$E y_{n+\tau} \overline{y_n} = E \xi e^{i(n+\tau)\omega} \overline{\xi} e^{-in\omega} = \sigma^2 e^{i\tau\omega}$$

is independent of n . \square

Exercise 2.4 Show that y is singular, i.e.

$$H_n^y = H_{n-1}^y \quad \text{for all } n.$$

Solution: A simple proof is obtained by applying the result given in the following exercise. \square

Exercise 2.5 Let y be a w.s.st. process such that $\dim(H_n^y) < \infty$ for some n . Then y is singular.

Solution: Let $\dim(H_n^y) = p < \infty$. Consider is a basis (y_{n-n_k}) , $k = 1, \dots, p$. This is a generator system, thus

$$y_n = \sum_{k=1}^p c_k y_{n-n_k},$$

with some $c_k \in \mathbb{R}$. It means, y_n can be predicted with zero error. \square

Exercise 2.6

$$y_n = \sum_{k=1}^m \xi_k e^{in\omega_k}$$

Prove that the process y defined above is singular, i.e.

$$H_n^y = H_{n-1}^y \quad \text{for all } n.$$

Solution: We apply previous Exercise. Let us consider $m + 1$ random variables in H_n^y : $y_{n_0}, y_{n_1}, \dots, y_{n_m}$. Then we can write

$$\begin{pmatrix} y_{n_0} \\ y_{n_1} \\ \vdots \\ y_{n_m} \end{pmatrix} = \begin{pmatrix} e^{in_0\omega_1} & e^{in_0\omega_2} & \dots & e^{in_0\omega_m} \\ e^{in_1\omega_1} & e^{in_1\omega_2} & \dots & e^{in_1\omega_m} \\ \vdots & \vdots & & \vdots \\ e^{in_m\omega_1} & e^{in_m\omega_2} & \dots & e^{in_m\omega_m} \end{pmatrix} \cdot \begin{pmatrix} \xi_1 \\ \xi_2 \\ \vdots \\ \xi_m \end{pmatrix},$$

or in compact form $Y = A \cdot Z$. Here Y and Z are random variables, and A is a matrix with constant elements from \mathbb{C} . The dimension of the matrix is $(m + 1) \times m$, thus the rows are linearly dependent. There is an $\alpha \in \mathbb{C}^{m+1}$ vector such that $\alpha^T A = 0$ and thus $\alpha^T Y = 0$. It means, that any $m + 1$ element in H_n^y is linearly dependent, thus $\dim(H_n^y) \leq m < \infty$, and the process is singular. \square

Exercise 2.7 A simple example for a real-valued singular process is given by

$$y_n = \cos(\omega n + \varphi) \quad \omega \neq 0,$$

where φ is a random phase with uniform distribution on $[0, 2\pi]$. Show that (y_n) is a wide sense stationary process.

Solution: As φ is a random variable with uniform distribution on $[0, 2\pi]$, its density function is

$$f(x) = \begin{cases} \frac{1}{2\pi} & x \in [0, 2\pi] \\ 0 & \text{otherwise} \end{cases}$$

Thus the mean value

$$\mathbb{E}y_n = \int_0^{2\pi} \cos(n\omega + x) \frac{1}{2\pi} dx = 0.$$

Applying the identity

$$\cos(\alpha) \cos(\beta) = \frac{\cos(\alpha - \beta) + \cos(\alpha + \beta)}{2}$$

we get

$$y_{n+\tau}y_n = \cos(\omega(n + \tau) + \varphi) \cos(\omega n + \varphi) = \frac{\cos(\omega\tau) + \cos(\omega(2n + \tau) + 2\varphi)}{2}.$$

Taking expectation we get using the above argument, that

$$\mathbb{E}y_{n+\tau}y_n = \frac{\cos(\omega\tau) + 0}{2},$$

which is indeed independent of n . \square

Exercise 2.8 Show that (y_n) is singular.

Solution: Apply the identity $\cos(\alpha + \beta) = \cos(\alpha)\cos(\beta) - \sin(\alpha)\sin(\beta)$. We get

$$y_{n+1} = \cos((\omega n + \varphi) + \omega) = \cos(\omega n + \varphi)\cos(\omega) - \sin(\omega n + \varphi)\sin(\omega),$$

thus

$$y_{n+1} = y_n \cos(\omega) \pm \sqrt{1 - y_n^2} \sin(\omega) \in H_n^y.$$

It means, that

$$H_n^y = H_{n-1}^y \quad \text{for all } n.$$

□

2.3 Wold decomposition

Lemma 1 For any random variable $\xi \in L_2(\Omega, \mathcal{F}, P)$ we have

$$\lim_{m \rightarrow \infty} (\xi | H_{-m}) = (\xi | H_{-\infty}).$$

Exercise 2.9 * Prove Lemma 1.

Solution: We shall apply the dual of this lemma, Lemma 2 below. Let us denote the orthogonal complement of H_{-m} in $L_2(\Omega, \mathcal{F}, P)$ by H_m^+ , i.e.

$$L_2(\Omega, \mathcal{F}, P) = H_{-m} \oplus H_m^+.$$

Then

$$(\xi | H_{-m}) = \xi - (\xi | H_m^+).$$

Applying Lemma 2 we get

$$\lim_{m \rightarrow \infty} (\xi | H_m^+) = (\xi | H_\infty^+), \quad H_\infty^+ = \text{cl} \left(\bigcup_m H_m^+ \right).$$

As H_∞^+ is the orthogonal complement of $H_{-\infty}$, Lemma 1 is proved. □

Lemma 2 Let $H_n \subset L_2(\Omega, \mathcal{F}, P)$ be a monotone increasing sequences of Hilbert subspaces, i.e. $H_n \subset H_{n+1}$. Let

$$H_\infty = \text{cl} \left(\bigcup_n H_n \right),$$

with cl denoting the closure. Then for any $x \in L_2(\Omega, \mathcal{F}, P)$ we have

$$\lim_{n \rightarrow \infty} (x | H_n) = (x | H_\infty).$$

Exercise 2.10 * *Prove Lemma 2.*

Solution: Let $(x|H_\infty) = y$, and write

$$x = y + \Delta x,$$

where $\Delta x \perp H_\infty$. Then also $\Delta x \perp H_n$ for any n . Thus projecting both sides of the above equality onto H_n we get

$$(x|H_n) = (y|H_n).$$

But $y \in \text{cl}(\cup H_n)$ implies that

$$\lim_{n \rightarrow \infty} (y - (y|H_n)) = 0,$$

thus $(y|H_n)$ converges to y , and the proposition follows. \square

Exercise 2.11 *Show that the processes (y_n^s) and (y_n^r) are orthogonal, $y^s \perp y^r$, meaning that*

$$y_n^s \perp y_m^r \quad \text{for all } n, m.$$

Solution: For any $v \in H_{-\infty}^y$ and any k we have $v \perp e_k$, since $v \in H_{k-1}^y$. By definition $y_n^s \in H_{-\infty}^y$ (since $y_n^s = (y_n|H_{-\infty}^y)$), and $y_n^r = \sum_{k=0}^{\infty} c_k e_{n-k}$. Thus for all n and m , y_n^s and y_m^r are indeed orthogonal. \square

Exercise 2.12 *Show that for the process $y^r = (y_n^r)$ we have*

$$H_{-\infty}^{y^r} = \{0\}.$$

Solution: First, by definition $y_n^r = \sum_{k=0}^{\infty} c_k e_{n-k}$, thus $H_n^{y^r} \subset H_n^e$ for all n . On the other hand $H_{-\infty}^e = \{0\}$, since $e_n \perp H_{-\infty}^e$ for all n . \square

Chapter 3

Spectral theory I.

3.1 The need for a spectral theory

Let

$$y_n = e_n + ce_{n-1}.$$

Then

$$e_n = -ce_{n-1} + y_n,$$

and iterating this equation we get, assuming $|c| < 1$,

$$e_n = \sum_{k=0}^{\infty} (-c)^k y_{n-k}.$$

Exercise 3.1 *Show that the right hand side above is well defined.*

Solution: Let us define

$$e_n^N = \sum_{k=0}^N (-c)^k y_{n-k}.$$

Then for $N < M$ we have

$$e_n^M - e_n^N = \sum_{k=N+1}^M (-c)^k y_{n-k},$$

thus

$$\mathbb{E}(e_n^M - e_n^N)^2 = r^y(0) \cdot \sum_{k=N+1}^M c^{2k} - 2r^y(1) \cdot \sum_{k=N+2}^M (c^{2k+1} + c^{2k-1}),$$

since $r^y(k) = 0$ if $|k| > 1$. It follows, that (e_n^M) is a Cauchy sequence as $M \rightarrow \infty$, thus the limit is well defined. \square

3.2 Fourier methods for w.s.st. processes

Proposition 1 *We have*

$$\begin{aligned}\lim_{N \rightarrow \infty} \xi_N(\omega_k) &= \xi_k, \\ \lim_{N \rightarrow \infty} \xi_N(\omega) &= 0 \quad \text{for } \omega \neq \omega_k\end{aligned}$$

in the sense of $L_2(\Omega, \mathcal{F}, P)$ and also w.p.1.

Exercise 3.2 *Prove Proposition 1.*

Solution:

$$\begin{aligned}\xi_N(\omega) &= \frac{1}{2N+1} \sum_{n=-N}^N y_n e^{-in\omega} = \frac{1}{2N+1} \sum_{n=-N}^N \left(\sum_{k=1}^m \xi_k e^{in\omega_k} \right) e^{-in\omega} = \\ &= \sum_{k=1}^m \xi_k \left(\frac{1}{2N+1} \sum_{n=-N}^N e^{in(\omega_k - \omega)} \right) = \sum_{k=1}^m c_k \xi_k.\end{aligned}$$

If $\omega = \omega_k$ for some k , then on the right hand side of the equation the coefficient of ξ_k is the following:

$$\frac{1}{2N+1} \sum_{n=-N}^N e^{in(\omega_k - \omega_k)} = \frac{1}{2N+1} \sum_{n=-N}^N 1 = 1,$$

thus this term of the sum is exactly ξ_k .

If $\omega \neq \omega_k$, the coefficient of ξ_k is the following:

$$c_k = \frac{1}{2N+1} \sum_{n=-N}^N e^{in(\omega_k - \omega)} = \frac{1 + 2 \sum_{n=1}^N \cos(n(\omega_k - \omega))}{2N+1}.$$

We shall use the formula for the Dirichlet kernel:

$$1 + 2 \sum_{k=1}^n \cos(kx) = \frac{\sin((n + \frac{1}{2})x)}{\sin(x/2)}.$$

Thus

$$c_k = \frac{1}{2N+1} \frac{\sin((N + \frac{1}{2})(\omega_k - \omega))}{\sin((\omega_k - \omega)/2)} \rightarrow 0 \quad \text{for } N \rightarrow \infty.$$

It proves the Proposition. □

Corrolary 1 *The spectral weights σ_k^2 can be obtained as*

$$\sigma_k^2 = E|\xi_k|^2 = \lim_{N \rightarrow \infty} E \left| \frac{1}{2N+1} \sum_{n=-N}^{+N} y_n e^{-in\omega_k} \right|^2.$$

Exercise 3.3 *Prove the above corollary.*

Solution: This follows simply from the fact that $\xi_N(\omega_k) \rightarrow \xi_k$ in $L_2(\Omega, \mathcal{F}, P)$. □

3.3 Herglotz's theorem

Exercise 3.4 *Show that the truncated $r_N(\tau)$ sequence itself is an auto-covariance sequence.*

Solution: Obviously, $r_N(\tau)$ is a positive semi-definite sequence. Hence it is an auto-covariance sequence.

Chapter 4

Spectral theory II.

4.2 Random orthogonal measures. Integration

Exercise 4.1 *Prove, that for any $0 \leq a < b < 2\pi$ we have*

$$F(b) - F(a) = \mathbb{E} |\zeta(b) - \zeta(a)|^2,$$

thus F is monotone nondecreasing.

Solution: Write $[0, b]$ as the union of $[0, a]$ and $[a, b]$. Then by definition

$$\zeta(b) - \zeta(a) \perp \zeta(a) - \zeta(0).$$

Apply Pythagoras theorem:

$$\mathbb{E} |\zeta(b) - \zeta(a)|^2 + \mathbb{E} |\zeta(a) - \zeta(0)|^2 = \mathbb{E} |\zeta(b) - \zeta(0)|^2.$$

Substituting $\zeta(0) = 0$, and $\mathbb{E} |\zeta(a)|^2 =: F(a)$ we get

$$\mathbb{E} |\zeta(b) - \zeta(a)|^2 + F(a) = F(b).$$

□

Exercise 4.2 *Let f, g be two left continuous step functions on $[0, 2\pi]$. Then*

$$\mathbb{E} I(f) \overline{I(g)} = \int_0^{2\pi} f(\omega) \overline{g(\omega)} dF(\omega). \quad (4.1)$$

Solution: Take a common subdivision for f and g , denote the intervals by $[\alpha_k, \beta_k)$, $k = 1, 2, \dots, n$. Then

$$\begin{aligned} f(\omega) &= \sum_{k=1}^n \lambda_k \chi_{[\alpha_k, \beta_k)}, \\ g(\omega) &= \sum_{k=1}^n \eta_k \chi_{[\alpha_k, \beta_k)}, \end{aligned}$$

and the product function:

$$f(\omega) \overline{g(\omega)} = \sum_{k=1}^n \lambda_k \overline{\eta_k} \chi_{[\alpha_k, \beta_k)}.$$

Thus the integral of $f\overline{g}$ on the right hand side of (4.1) is the following:

$$\int_0^{2\pi} f(\omega) \overline{g(\omega)} dF(\omega) = \sum_{k=1}^n \lambda_k \overline{\eta_k} (F(\beta_k) - F(\alpha_k)).$$

Let us compute the left hand side of (4.1). We have

$$I(f) = \sum_{k=1}^n \lambda_k (\zeta(\beta_k) - \zeta(\alpha_k)), \quad I(g) = \sum_{k=1}^n \eta_k (\zeta(\beta_k) - \zeta(\alpha_k)).$$

The random variables $(\zeta(\beta_k) - \zeta(\alpha_k))$ $k = 1, 2, \dots, n$ are independent, thus

$$EI(f) \overline{I(g)} = \sum_{k=1}^n \lambda_k \overline{\eta_k} E|\zeta(\beta_k) - \zeta(\alpha_k)|^2 = \sum_{k=1}^n \lambda_k \overline{\eta_k} (F(\beta_k) - F(\alpha_k)).$$

□

Exercise 4.3 Prove that

$$y_n = \int_0^{2\pi} e^{in\omega} d\zeta(\omega)$$

is w.s.st.

Solution: $Ey_n = 0$. Applying the isometry we get

$$Ey_n \overline{y_{n+\tau}} = \int_0^{2\pi} e^{in\omega} \overline{e^{i(n+\tau)\omega}} dF(\omega) = \int_0^{2\pi} e^{i(-\tau)\omega} dF(\omega),$$

which is independent of n .

□

4.3 Representation of a wide sense stationary process

Exercise 4.4 *Prove the above implication.*

Solution: Let $g(\omega) = \sum_{n \in \mathbb{N}} c_n e^{in\omega}$, $c_n \in \mathbb{C}$. Then $g = 0$ in $L_2^c(dF) = L_2^c([0, 2\pi), dF)$ means, that

$$\int_0^{2\pi} |g(\omega)|^2 dF(\omega) = 0.$$

The corresponding random variable is $I(g) = \sum_{n \in \mathbb{N}} c_n y_n$. Then

$$\mathbb{E}I(g) = 0, \quad \mathbb{E}I(g)\overline{I(g)} = \sum_n \sum_m c_n \overline{c_m} \int_0^{2\pi} e^{i(n-m)\omega} dF(\omega) = \int_0^{2\pi} |g(\omega)|^2 dF(\omega) = 0.$$

Thus $I(g) = 0$ a.s. in $L_2^c(\Omega, \mathcal{F}, P)$. □

Exercise 4.5 *Show that the structure function of the random orthogonal measure ζ is F , predetermined by the spectral distribution function of y .*

Solution: Indeed, we have

$$dF(\omega) = dF^y(\omega) \cdot \frac{1}{2\pi},$$

since $y_n = \int_0^{2\pi} e^{in\omega} d\zeta(\omega)$, thus

$$\mathbb{E}(y_{n+\tau} \overline{y_n}) = \int_0^{2\pi} e^{i(n+\tau)\omega} e^{-in\omega} dF(\omega) = \int_0^{2\pi} e^{i\tau\omega} dF(\omega).$$

□

4.4 Change of measure

Let $d\zeta(\omega)$ be a random orthogonal measure on $[0, 2\pi)$ with the structure function $F(\omega)$. Let $g \in L_2^c(dF)$, and define

$$\eta(\omega) = \int_0^\omega g(\omega') d\zeta(\omega') \quad 0 \leq \omega < 2\pi.$$

Exercise 4.6 *Show that $d\eta(\omega)$ is a random orthogonal measure, with the structure function*

$$dG(\omega) = |g(\omega)|^2 dF(\omega).$$

Solution: We have $\eta(0) = 0$. For any two non-overlapping intervals $[a, b)$ and $[c, d)$ contained in $[0, 2\pi)$ we have

$$\eta(d) - \eta(c) = \int_c^d g(\omega') d\zeta(\omega'), \quad \eta(b) - \eta(a) = \int_a^b g(\omega') d\zeta(\omega').$$

As the domains of integration are non-overlapping,

$$\eta(d) - \eta(c) \perp \eta(b) - \eta(a).$$

The structure function can be computed using the isometry between $L_2^c(dF)$ and $L_2^c(\Omega, \mathcal{F}, P)$:

$$\mathbb{E}|\eta(a)|^2 = \mathbb{E}\eta(a)\overline{\eta(a)} = \int_0^{2\pi} g(\omega)\overline{g(\omega)}dF(\omega) = \int_0^{2\pi} |g(\omega)|^2 dF(\omega).$$

□

4.5 Linear filters

Exercise 4.7 Show that the spectral representation process of y is given by

$$d\zeta^y(\omega) = H(e^{-i\omega})d\zeta^u(\omega),$$

where

$$H(e^{-i\omega}) = \sum_{k=0}^m h_k e^{-ik\omega}.$$

Solution: The spectral representation of (u_n) gives

$$u_n = \int_0^{2\pi} e^{in\omega} d\zeta^u(\omega).$$

Thus

$$u_{n-k} = \int_0^{2\pi} e^{i(n-k)\omega} d\zeta^u(\omega), \quad \text{and} \quad h_k u_{n-k} = \int_0^{2\pi} h_k e^{i(n-k)\omega} d\zeta^u(\omega).$$

It follows, that

$$\begin{aligned} y_n &= \sum_{k=0}^m h_k u_{n-k} = \int_0^{2\pi} \sum_{k=0}^m h_k e^{i(n-k)\omega} d\zeta^u(\omega) = \\ &= \int_0^{2\pi} e^{in\omega} \left(\sum_{k=0}^m h_k e^{-ik\omega} d\zeta^u(\omega) \right) = \int_0^{2\pi} e^{in\omega} d\zeta^y(\omega). \end{aligned}$$

□

Exercise 4.8 *Show that the integrand on the right hand side converges to $e^{in\omega}H(e^{-i\omega})$ in $L_2^c(dF^u)$.*

Solution: We know, that

$$\lim_{m \rightarrow \infty} H^m(e^{-i\omega}) = \lim_{m \rightarrow \infty} \sum_{k=0}^m h_k e^{-ik\omega} = y_n$$

in $L_2^c(\Omega, \mathcal{F}, P)$. Thus due to isometry

$$\lim_{m \rightarrow \infty} \int_0^{2\pi} e^{in\omega} H^m(e^{-i\omega}) d\zeta^u(\omega) = \int_0^{2\pi} e^{in\omega} H(e^{-i\omega}) d\zeta^u(\omega).$$

□

Chapter 5

AR, MA and ARMA processes

5.3 The AR(1) process

Exercise 5.1 Prove that (y_n) defined by $y_n = \sum_{k=0}^{\infty} (-a)^k e_{n-k}$ does indeed satisfy $y_n + ay_{n-1} = e_n$.

Solution: We have

$$\begin{aligned} y_n &= e_n + \sum_{k=1}^{\infty} (-a)^k e_{n-k} = e_n + (-a) \sum_{k=1}^{\infty} (-a)^{k-1} e_{n-1-(k-1)} = \\ &= e_n + (-a) \sum_{k=0}^{\infty} (-a)^k e_{n-1-k} = e_n - ay_{n-1}. \end{aligned}$$

Here we used the fact, that the coefficient-sequence is absolute convergent, and (e_{n-k}) is an orthogonal sequence in $L_2(\Omega, \mathcal{F}, P)$. \square

Exercise 5.2 If $|a| > 1$ define

$$y_n = \sum_{k=0}^{\infty} \left(-\frac{1}{a}\right)^k \frac{1}{a} e_{n+k+1}.$$

Show that the r.h.s. is well-defined, and (y_n) does indeed satisfy $y_n + ay_{n-1} = e_n$.

Solution: The r.v-s (e_{n+k}) , $k = 1, 2, \dots$ form an orthogonal sequence in $L_2(\Omega, \mathcal{F}, P)$. As $|a| > 1$, the coefficient sequence on the right hand side is absolute convergent. Thus the right hand side is well-defined.

We have

$$y_n = \frac{1}{a}e_{n+1} + \frac{1}{a} \sum_{k=1}^{\infty} \left(-\frac{1}{a}\right)^{k-1} \left(-\frac{1}{a}\right) e_{n+k+1} = \frac{1}{a}e_{n+1} - \frac{1}{a} \sum_{k=0}^{\infty} \left(-\frac{1}{a}\right)^k \frac{1}{a} e_{(n+1)+(k+1)},$$

thus

$$y_n = \frac{1}{a}e_{n+1} - \frac{1}{a}y_{n+1},$$

and it can be rearranged as

$$e_{n+1} = y_{n+1} + ay_n.$$

□

5.5 MA processes

Thus let $e = (e_n)$ be a wide sense stationary orthogonal process and define

$$y_n = \sum_{k=0}^m c_k e_{n-k}$$

Exercise 5.3 Show that (y_n) is a wide sense stationary process.

Solution: The mean value and the covariances are independent of n , since $Ey_n = 0$, and

$$E(y_n^2) = \sum_{k=0}^m \sum_{j=0}^m c_k c_j E(e_{n-k} e_{n-j}) = \sum_{k=0}^m c_k^2 \sigma^2.$$

For $0 < \tau \leq m$ we have:

$$E(y_{n+\tau} y_n) = \sum_{k=0}^m \sum_{j=0}^m c_k c_j E(e_{n-k+\tau} e_{n-j}) = \sum_{k=\tau}^m c_k c_{k-\tau}.$$

Similarly, for $\tau > m$ we get $E(y_{n+\tau} y_n) = 0$.

□

Exercise 5.4 Show that $e_n = \sum_{k=0}^{\infty} (-1)^k b^k y_{n-k}$ is well-defined, if $\sum_{k=0}^{\infty} b^{2k} < \infty$, i.e. $|b| < 1$. Similarly, $e_n = \sum_{k=0}^{\infty} (-1)^k \frac{1}{b^k} z_{n-k}$ is well-defined if $|b| > 1$.

Solution: We shall use the Cauchy-Schwartz-Buniakovskii inequality in $L_2(\Omega, \mathcal{F}, P)$. Thus

$$|E(y_n y_{n-\tau})| = |\langle y_n, y_{n-\tau} \rangle| \leq \|y_n\| \|y_{n-\tau}\| = \sqrt{Ey_n^2 Ey_{n-\tau}^2} = r^y(0).$$

Let us consider the $|b| < 1$ case. We have

$$e_n = \lim_{N \rightarrow \infty} e_n^N, \quad \text{with} \quad e_n^N = \sum_{k=0}^N (-1)^k b^k y_{n-k}.$$

Then using CSB we get

$$E(e_n^N)^2 \leq \sum_{k=0}^N b^{2k} r^y(0) < r^y(0) \frac{1}{1-b^2}.$$

Thus (e_n^N) , $N \in \mathbb{N}$ is a uniformly convergent Cauchy sequence, it is convergent. \square

5.6 ARMA processes

Proposition 2 *Assume that $A(z^{-1}) \neq 0$ for $|z| = 1$. Then there is a unique w.s.st. process satisfying $A(q^{-1})y = C(q^{-1})e$.*

Exercise 5.5 *Prove the above proposition.*

Solution: Let us assume, that such a process does exist. As (y_n) is a wide sense stationary process, there is a unique random orthogonal measure $d\zeta^y(\omega)$, such that

$$y_n = \int_0^{2\pi} e^{in\omega} d\zeta^y(\omega).$$

The spectral representation of the process $u_n = A(q^{-1})y_n = C(q^{-1})e_n$ is

$$u_n = \int_0^{2\pi} A(e^{-in\omega}) d\zeta^y(\omega) = \int_0^{2\pi} C(e^{-in\omega}) d\zeta^e(\omega).$$

From the uniqueness of the spectral representation it follows, that

$$d\zeta^y(\omega) = \frac{C(e^{-in\omega})}{A(e^{-in\omega})} d\zeta^e(\omega).$$

As (e_n) has spectral density function, (y_n) also has spectral density function::

$$f^y(\omega) = \frac{C(e^{in\omega})C(e^{-in\omega})}{A(e^{in\omega})A(e^{-in\omega})} f^y(\omega) = \frac{1}{2\pi} |H(e^{-in\omega})|^2,$$

where

$$H(e^{-in\omega}) = \frac{C(e^{-in\omega})}{A(e^{-in\omega})}.$$

Now, since $A(-e^{in\omega}) \neq 0$ for $\omega \in [0, 2\pi]$, obviously $f(\omega) \geq 0$. Thus there exist a process with this spectral density function. \square

Proposition 3 *Assume that $A(z^{-1})$ and $C(z^{-1})$ are stable polynomials. Then (e_n) is the innovation process of (y_n) .*

Exercise 5.6 *Prove the above proposition.*

5.8 ARMA processes with unstable zeros

Exercise 5.7 Let G be an all-pass transfer function, and let e' be a w.s.st. orthogonal process. Then the process e defined by

$$d\zeta^e(\omega) = G(e^{-i\omega})d\zeta^{e'}(\omega)$$

is also a w.s.st. orthogonal process.

Solution: The spectral density of e is

$$f(\omega) = |G(e^{-i\omega})|^2 f_{e'}(\omega) = f_{e'}(\omega),$$

since $|G(e^{-i\omega})|^2 = 1$. Thus the spectral density functions of e and e' are the same, e' is also a w.s.st. orthogonal process \square

Chapter 6

Multivariate time series

6.1 Vector valued wide sense stationary processes

Exercise 6.1 Show that (z_n) is an \mathbb{R} -valued wide sense stationary process and we have

$$\mathbb{E} z_n^2 = \sum_{k=1}^p \sum_{l=1}^p a_k^\top R(k-l) a_l \geq 0. \quad (6.1)$$

Solution:

$$\begin{aligned} \mathbb{E} z_n^2 &= \mathbb{E} \sum_{k=1}^p a_k^\top y_{n-k} \left(\sum_{l=1}^p a_l^\top y_{n-l} \right)^\top = \mathbb{E} \sum_{k=1}^p a_k^\top y_{n-k} \cdot \sum_{l=1}^p y_{n-l}^\top a_l = \\ &= \sum_{k=1}^p \sum_{l=1}^p a_k^\top (\mathbb{E} y_{n-k} y_{n-l}^\top) a_l = \sum_{k=1}^p \sum_{l=1}^p a_k^\top (R(n-k - (n-l))) a_l. \end{aligned}$$

□

Exercise 6.2 Prove the converse statement: let $R(\tau)$, $-\infty < \tau < +\infty$ be a positive definite sequence of real-valued, $s \times s$ matrices. Then it is the auto-covariance sequence of an \mathbb{R}^s -valued, wide sense stationary Gaussian process.

Solution: (Sketch) First truncate the sequence as

$$R_N(\tau) = \begin{cases} R(\tau) & \text{if } |\tau| \leq N \\ 0 & \text{if } |\tau| > N \end{cases}$$

Then there are random variables $\xi_{-N}, \dots, \xi_0, \dots, \xi_N$ with Gaussian distribution, such that the autocovariances are the given $R_N(\tau)$.

Now we can apply Kolmogorov extension theorem, since the measures satisfy the consistency conditions as $N \rightarrow \infty$. This theorem ensures, that there exists a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and a stochastic process with the above marginal distributions. □

6.2 Prediction and the innovation process

Exercise 6.3 Show that $e_n \equiv 0$ is indeed equivalent to

$$H_n^y = H_{n-1}^y \text{ for all } n.$$

Solution: It is the same, as in the scalar case.

6.3 Spectral theory

Exercise 6.4 Prove that a quadratic form $\alpha^\top F \alpha$, with F symmetric, determines the bilinear form corresponding to F uniquely as

$$\beta^\top F \gamma = \frac{1}{4} ((\beta + \gamma)^\top F (\beta + \gamma) - (\beta - \gamma)^\top F (\beta - \gamma)).$$

Solution: We are in a Euclidian space. This formula gives the well-known connection between the norm and the scalar product. Especially, for $\beta = \gamma$ we get back the quadratic form:

$$\beta^\top F \beta = \frac{1}{4} ((2\beta)^\top F (2\beta) - (\beta - \beta)^\top F (\beta - \beta)) = \frac{1}{4} ((2\beta)^\top F (2\beta)).$$

Exercise 6.5 * Let $F(\cdot)$ be an $s \times s$ matrix-valued function on $[0, 2\pi]$ such that the increments $F(\cdot)$ are symmetric and positive semidefinite. Then for any $k, l = 1, \dots, s$ the elements $F_{k,l}(\omega)$ are of finite variations.

Solution: For any fixed $\alpha \in \mathbb{R}^s$ the function $F_\alpha(\omega) = \alpha^\top F(\omega) \alpha$ is monotone non-decreasing. Recall the expression of the bilinear form $\beta^\top F(\omega) \gamma$ via the corresponding quadratic form.

Exercise 6.6 Show that for an \mathbb{R}^s -valued orthogonal wide sense stationary process (e_n) with covariance matrix $\Lambda = \mathbb{E} e_n e_n^\top$ we have

$$f(\omega) = \Lambda \quad \text{for } \forall \omega.$$

Solution: Using the given $f(\omega)$ function we get

$$\frac{1}{2\pi} \int_0^{2\pi} e^{i\tau\omega} f(\omega) d\omega = \Lambda \frac{1}{2\pi} \int_0^{2\pi} e^{i\tau\omega} d\omega = \begin{cases} \Lambda & \text{if } \tau = 0 \\ 0 & \text{if } \tau \neq 0 \end{cases},$$

and this is the autocovariance matrix function of an orthogonal wide sense stationary process with covariance matrix Λ . \square

6.4 Filtering

Proposition 4 *The spectral distribution of the process v is given by*

$$dF^v(\omega) = H(e^{-i\omega})dF^y(\omega)H(e^{i\omega})^\top.$$

Exercise 6.7 *Prove Proposition 4.*

Solution: We have

$$E y_{n+\tau} y_n^T = R(\tau) = \frac{1}{2\pi} \int_0^{2\pi} e^{i\tau\omega} dF(\omega).$$

Then

$$\begin{aligned} E v_{n+\tau} v_n^T &= E \sum_{k=0}^p h_k y_{n+\tau-k} \left(\sum_{j=0}^p h_j y_{n-j} \right)^T = \\ &= \sum_{k=0}^p \sum_{j=0}^p h_k E (y_{n+\tau-k} y_{n-j}^T) h_j^T = \sum_{k=0}^p \sum_{j=0}^p h_k \left(\frac{1}{2\pi} \int_0^{2\pi} e^{i(\tau-k+j)\omega} dF(\omega) \right) h_j^T \\ &= \frac{1}{2\pi} \int_0^{2\pi} e^{i\tau\omega} \left(\sum_{k=0}^p h_k e^{-ik\omega} \right) dF(\omega) \left(\sum_{j=0}^p h_j^T e^{ij\omega} \right) = \frac{1}{2\pi} \int_0^{2\pi} e^{i\tau\omega} dF^v(\omega). \end{aligned}$$

□

Proposition 5 *Assume that the right hand side of*

$$H(e^{-i\omega}) = \sum_{k=0}^{\infty} h_k e^{-i\omega k}$$

converges in $L_2^{c,r \times s}(dF)$. Then the process (v_n) under

$$v_n = \sum_{k=0}^{\infty} h_k y_{n-k}$$

is well-defined, i.e. the right hand side converges in $L_2^{c,r}(\Omega, \mathcal{F}, P)$, and the spectral distribution of (v_n) is given by

$$dF^v(\omega) = H(e^{-i\omega})dF^y(\omega)H^\top(e^{i\omega}).$$

Exercise 6.8 *Prove the above proposition following the proof for the scalar case.*

Solution: Truncate the infinite sum defining v_n at m , i.e. define

$$v_n^m = \sum_{k=0}^m h_k y_{n-k}.$$

Then the spectral representation of $v^m = (v_n^m)$ is given as

$$v_n^m = \int_0^{2\pi} e^{in\omega} H^m(e^{-i\omega}) d\zeta^u(\omega),$$

where

$$H^m(e^{-i\omega}) = \sum_{k=0}^m h_k e^{-ik\omega}.$$

Now letting m tend to infinity, the l.h.s. of the above equality converges to v_n in $L_2^{c,r}(\Omega, \mathcal{F}, P)$. Similarly, the integrand on the right hand side converges to $e^{in\omega} H(e^{-i\omega})$ in $L_2^{c,r \times s}(dF)$ \square

6.5 Multivariate random orthogonal measures

Theorem 1 *Let $d\zeta(\omega)$ be a \mathbb{C}^s -valued random orthogonal measure, with structure function $dF(\omega)$. Then the \mathbb{C}^s -valued process*

$$y_n = \int_0^{2\pi} e^{in\omega} d\zeta(\omega)$$

is wide sense stationary, and its spectral distribution function is given by

$$dF^y(\omega) = 2\pi dF(\omega).$$

Exercise 6.9 *Prove Theorem 1.*

Solution: Apply previous Theorem about the isometry, stating

$$\mathbb{E} I(g) I(h)^* = \int_0^{2\pi} g(\omega) \bar{h}(\omega) dF(\omega)$$

where

$$I(g) = \int_0^{2\pi} g(\omega) \cdot I d\zeta(\omega), \quad I(h) = \int_0^{2\pi} h(\omega) \cdot I d\zeta(\omega).$$

Then

$$y_n = I(e^{in\omega}), \quad y_{n+\tau} = I(e^{i(n+\tau)\omega}).$$

Thus

$$\begin{aligned} \mathbb{E} y_n y_{n+\tau}^* &= \mathbb{E} I((e^{in\omega}) I(e^{i(n+\tau)\omega})^*) = \int_0^{2\pi} e^{in\omega} \bar{e}^{i(n+\tau)\omega} dF(\omega) = \\ &= \int_0^{2\pi} e^{-i\tau\omega} dF(\omega) = \frac{1}{2\pi} \int_0^{2\pi} e^{-i\tau\omega} 2\pi dF(\omega) = \frac{1}{2\pi} \int_0^{2\pi} e^{-i\tau\omega} dF^y(\omega). \end{aligned}$$

\square

6.7 Linear filters

Exercise 6.10 Show that the spectral representation process of v is given by

$$d\zeta^v(\omega) = H(e^{-i\omega})d\zeta^y(\omega).$$

Solution: The spectral representation of (y_n) gives

$$y_n = \int_0^{2\pi} e^{in\omega} d\zeta^y(\omega).$$

Thus

$$y_{n-k} = \int_0^{2\pi} e^{i(n-k)\omega} d\zeta^y(\omega), \quad \text{and} \quad h_k y_{n-k} = \int_0^{2\pi} h_k e^{i(n-k)\omega} d\zeta^y(\omega).$$

It follows, that

$$\begin{aligned} v_n &= \sum_{k=0}^p h_k y_{n-k} = \int_0^{2\pi} \sum_{k=0}^p h_k e^{i(n-k)\omega} d\zeta^y(\omega) = \\ &= \int_0^{2\pi} e^{in\omega} \left(\sum_{k=0}^p h_k e^{-ik\omega} d\zeta^y(\omega) \right) = \int_0^{2\pi} e^{in\omega} d\zeta^v(\omega). \end{aligned}$$

□

Exercise 6.11 Assume, that the infinite series

$$H(e^{-i\omega}) = \sum_{k=0}^{\infty} h_k e^{-ik\omega}.$$

converges in $L_2^{c,r \times s}(dF^y)$. Then the spectral representation process of (v_n) is

$$d\zeta^v(\omega) = H(e^{-i\omega})d\zeta^y(\omega).$$

Solution: Truncate the infinite sum defining v_n at p , i.e. define

$$v_n^p = \sum_{k=0}^p h_k y_{n-k}.$$

Then the spectral representation of $v^p = (v_n^p)$ is given as

$$v_n^p = \int_0^{2\pi} e^{in\omega} H^p(e^{-i\omega}) d\zeta^u(\omega),$$

where

$$H^p(e^{-i\omega}) = \sum_{k=0}^p h_k e^{-ik\omega}.$$

Now letting p tend to infinity, the l.h.s. of the previous equality converges to v_n in $L_2^{c,s}(\Omega, \mathcal{F}, P)$. □

Exercise 6.12 *Re-derive the formula for the spectral distribution measure of v :*

$$dF^v(\omega) = H(e^{-i\omega})dF^y(\omega)H^\top(e^{i\omega})$$

using the exercise above.

Solution: The spectral representation process of (v_n) is

$$d\zeta^v(\omega) = H(e^{-i\omega})d\zeta^y(\omega),$$

which means that the process can be represented as

$$v_n = \int_0^{2\pi} e^{in\omega} d\zeta^v(\omega) = \int_0^{2\pi} e^{in\omega} H(e^{-i\omega}) d\zeta^y(\omega).$$

Then also we have

$$v_{n-\tau} = \int_0^{2\pi} e^{i(n-\tau)\omega} H(e^{-i\omega}) d\zeta^y(\omega).$$

The covariance of these two random vectors can be written using the isometry:

$$Ev_n v_{n-\tau}^T = \int_0^{2\pi} e^{in\omega} H(e^{-i\omega}) dF^y(\omega) \overline{e^{i(n-\tau)\omega} H(e^{-i\omega})} = \int_0^{2\pi} e^{i\tau\omega} H(e^{-i\omega}) dF^y(\omega) H(e^{i\omega}).$$

Comparing it with the definition of the spectral distribution measure:

$$Ev_n v_{n-\tau}^T = \int_0^{2\pi} e^{i\tau\omega} dF^v(\omega)$$

we get the desired formula. □

6.8 Proof of the spectral representation theorem

Exercise 6.13 *Let $\zeta(a), 0 \leq a < 2\pi$ be an \mathbb{C}^s -valued stochastic process such that for any $\alpha \in \mathbb{C}^s$ the complex-valued process $\alpha^\top \zeta(a), 0 \leq a < 2\pi$ has orthogonal increments. Then the process $\zeta(\cdot)$ itself has orthogonal increments.*

Solution: Indirect. Let us assume, that there exist $0 \leq a < b \leq c < d \leq 2\pi$, such that and the covariance matrix of the increments is not 0, i.e.

$$E(\zeta(b) - \zeta(a))(\overline{\zeta(d) - \zeta(c)})^\top = R \neq 0 \in \mathbb{R}^{s \times s}.$$

But on the other hand for any $\alpha \in \mathbb{R}^s$ $\alpha^* R \alpha = 0$. Then also $\alpha^* R^* \alpha = 0$, and thus the quadratic form

$$\alpha^* (R + R^*) \alpha = 0, \quad \implies \quad (R + R^*) = 0.$$

It means, that R is antisymmetric. □

Chapter 7

State-space representation

7.1 From multivariate AR(1) to state-space equations

Proposition 6 *Let us consider the multivariate linear stochastic equation*

$$x_{n+1} = Ax_n + Bv_n.$$

Let A be a stable $s \times s$ matrix. Then this equation has a unique wide sense stationary solution (x_n) , given by

$$x_{n+1} = \sum_{k=0}^{\infty} A^k B v_{n-k}.$$

It follows that x is a causal linear function of v , more exactly, for all n

$$H_{n+1}^x \subset H_n^v.$$

Exercise 7.1 *Prove the above proposition.*

Solution: The spectral representation of x is given by

$$d\zeta^x(\omega) = A(e^{-i\omega})^{-1} d\zeta^v(\omega),$$

where

$$A(e^{-i\omega})^{-1} = (e^{i\omega}I - A)^{-1} B.$$

As A is stable, $\det(zI - A)$ is not zero for $|z| \geq 1$. It follows, that

$$(zI - A)^{-1} = \sum_{k=0}^{\infty} A^k z^{-k}$$

is well defined for $|z| \geq 1$, i.e.

$$\sum_{k=0}^{\infty} A^k z^{-k}$$

is well defined for $|z| < 1$. Then

$$\sum_{k=0}^{\infty} A^k B v_{n-k}$$

is well defined. □

Lemma 3 *Let A be an $s \times s$ stable real matrix. Then*

$$(e^{i\omega} I - A)^{-1} = \sum_{k=1}^{\infty} h_k e^{-ik\omega}, \quad h_1 = I$$

*with some sequences of $s \times s$ real matrices h_k , where convergence in the right hand side is understood in the sense of $L_2^{c,s \times s}(d\omega)$. In fact, convergence is also **uniform** in ω .*

Exercise 7.2 *Prove the Lemma 3.*

Solution: Note that the matrix-valued function $(zI - A)^{-1}$ is analytic in $\{z : |z| > 1 - \varepsilon\}$ for some $\varepsilon > 0$. Then it is analytic for $|z| = 1$, it can be represented as a Taylor series for $z = e^{i\omega}$.

Exercise 7.3 *Prove the Proposition 6 using Lemma 3.*

Solution: It was already done in Exercise 7.1.

Proposition 7 *Assume, that $e^{i\omega} I - A$ is not singular for all $\omega \in [0, 2\pi]$. Then*

$$x_{n+1} = Ax_n + Bv_n$$

has a unique solution.

Exercise 7.4 *Prove the above Proposition 7.*

Solution: Assume that a wide sense stationary solution (x_n) exists. There is a unique random orthogonal measure $d\zeta^x(\omega)$, such that

$$x_n = \int_0^{2\pi} e^{in\omega} d\zeta^x(\omega).$$

Then also show that necessarily

$$x_{n+1} = \int_0^{2\pi} e^{i(n+1)\omega} d\zeta^x(\omega).$$

and

$$Ax_n + Bv_n = \int_0^{2\pi} e^{in\omega} (Ad\zeta^x(\omega) + B) d\zeta^v(\omega).$$

From the uniqueness of the orthogonal measure we get

$$d\zeta^x(\omega) = (e^{i\omega}I - A)^{-1}B d\zeta^v(\omega).$$

Then the spectral representation of the autocovariance matrices gives

$$dF^x(\omega) = (e^{i\omega}I - A)^{-1}BdF^y(\omega)((e^{-i\omega}I - A)^{-1}B)^\top.$$

The condition of Proposition 7 ensures, that $dF^x(\omega)$ is well-defined. Thus the state-space equation has a unique solution. \square

7.2 Auto-covariances and the Lyapunov-equation

Exercise 7.5 Show that P can be written as

$$P = \sum_{k=0}^{\infty} A^k B \Sigma_{vv} B^T (A^T)^k.$$

Solution: We can use the expression of (x_n) given by

$$x_{n+1} = \sum_{k=0}^{\infty} A^k B v_{n-k}.$$

The solution follows from the orthogonality of (v_n) , i.e.

$$E v_n^T v_m = \begin{cases} \Sigma_{vv} & \text{if } n = m \\ 0 & \text{if } n \neq m \end{cases}$$

\square

Exercise 7.6 Show directly, with purely algebraic arguments, that, if A is stable, the Lyapunov-equation

$$P = APA^T + B\Sigma_{vv}B^T$$

has a unique solution P , and show that it can be written in the form

$$\sum_{k=0}^{\infty} A^k B \Sigma_{vv} B^T (A^T)^k.$$

Prove that the solution P , given by this formula, is positive semi-definite.

Solution: Iterate equation

$$P = APA^T + B\Sigma_{vv}B^T.$$

We get

$$\begin{aligned} P &= APA^T + B\Sigma_{vv}B^T = A^2P(A^T)^2 + AB\Sigma_{vv}B^TA^T + B\Sigma_{vv}B^T = \dots \\ &= A^{n+1}P(A^T)^{n+1} + \sum_{k=0}^n A^k B\Sigma_{vv}B^T(A^T)^k. \end{aligned}$$

The first term on the last expression tends to zero with exponential rate, since A is stable. Thus we get the given representation. \square

Exercise 7.7 Consider two Lyapunov equations

$$P = APA^T + B\Sigma_{vv}B^T$$

with a **common** stable A and such that

$$B_1\Sigma_{1,vv}B_1^T \leq B_2\Sigma_{2,vv}B_2^T.$$

Let the solutions be denoted by P_1 and P_2 . Show that $P_1 \leq P_2$.

Solution: If A is stable, then

$$B_1\Sigma_{1,vv}B_1^T \leq B_2\Sigma_{2,vv}B_2^T \quad \implies \quad AB_1\Sigma_{1,vv}B_1^TA^T \leq AB_2\Sigma_{2,vv}B_2^TA^T.$$

Thus

$$P_1 = \sum_{k=0}^{\infty} A^k B_1\Sigma_{1,vv}B_1^T(A^T)^k \leq \sum_{k=0}^{\infty} A^k B_2\Sigma_{2,vv}B_2^T(A^T)^k = P_2.$$

Exercise 7.8 Prove that for $\tau < 0$ we have $R(\tau) = R(-\tau)^T$.

Solution:

$$R(\tau) = \mathbb{E}y_{n+\tau}y_n^T = (\mathbb{E}y_n y_{n+\tau}^T)^T = (\mathbb{E}y_{n-\tau}y_n^T)^T = R(-\tau)^T.$$

\square

Exercise 7.9 Prove the validity of the recursion

$$P_{n+1} = AP_nA^T + BB^T$$

for P_n .

To get the covariance function of (x_n) take the dyadic product of

$$x_{n+1} = Ax_n + Bv_n$$

with itself:

$$x_{n+1}x_{n+1}^T = Ax_nx_n^TA^T + Bv_nv_n^TB^T + Ax_nv_n^TB^T + Bv_nx_n^TA^T.$$

Now as $x_n \in H_{n-1}^v$, we have

$$x_n \perp v_n.$$

Then taking expectation on both sides of the above equation we get given recursion. \square

Exercise 7.10 *Show that if A is stable, then P_n converges to the unique solution of*

$$P = APA^T + B\Sigma_{vv}B^T.$$

Solution: Reiterating the previous recursion, we get

$$\begin{aligned} P_{n+1} &= AP_nA^T + BB^T = A(AP_{n-1}A^T + BB^T)A^T + BB^T = \\ &= A^2P_{n-1}A^{2T} + ABB^TA^T + BB^T = \dots = A^nP_0(A^n)^T + \sum_{k=0}^{n-1} A^kBB^T(A^T)^k. \end{aligned}$$

As A is stable:

$$\lim_{n \rightarrow \infty} A^nP_0(A^n)^T = 0,$$

thus

$$\lim_{n \rightarrow \infty} P_n = \sum_{k=0}^{\infty} A^kBB^T(A^T)^k = P_{\infty}.$$

In Exercise 7.6 we have shown that this is the solution of the discrete Lyapunov equation.

\square

7.3 State space representation of ARMA processes

Exercise 7.11 *Prove that R_N is non-singular by taking a state-space representation of y .*

Chapter 8

Kalman filtering

8.1 The filtering problem

Exercise 8.1 *Provide an argument for the validity of*

$$H_n^y = H_{n-1}^y \oplus \mathcal{L}(\nu_n).$$

Solution: We have $\nu_n = y_n - (y_n|H_{n-1}^y)$, thus ν_n is the innovation process of y_n . \square

8.2 The Kalman-gain matrix

Exercise 8.2 *Prove that the projection of the random vector $x \in L_2^s(\Omega, \mathcal{F}, P)$ onto the finite dimensional subspace of $L_2(\Omega, \mathcal{F}, P)$ spanned by the components of ν is given by*

$$\hat{x} = E(x\nu^T)(E(\nu\nu^T))^{-1}\nu.$$

Solution: Obviously, \hat{x} is an element of the finite dimensional subspace of $L_2(\Omega, \mathcal{F}, P)$ spanned by the components of ν . Also, $x - \hat{x}$ is orthogonal to the components of ν , since

$$\begin{aligned} E((x - \hat{x})\nu^T) &= E x \nu^T - E \hat{x} \nu^T = E x \nu^T - E (E(x\nu^T)(E(\nu\nu^T))^{-1}\nu) \nu^T = \\ &= E x \nu^T - E(x\nu^T)(E(\nu\nu^T))^{-1}E(\nu\nu^T) = 0. \end{aligned}$$

Thus \hat{x} is the projection. \square

Exercise 8.3 *Derive the above expression*

$$H^{-1}(q^{-1}) = I - C(qI - A + KC)^{-1}K$$

from $H(q^{-1}) = C(qI - A)^{-1}K + I$ using the matrix inversion lemma.

Solution: THE MATRIX INVERSION LEMMA is the following: Let

$$F = \begin{pmatrix} A_0 & B_0 \\ C_0 & D_0 \end{pmatrix}$$

be a 2×2 block-matrix with A_0 and D_0 being square matrices. Assume that A_0 , D_0 are non-singular and so is $A_0 - B_0D_0^{-1}C_0$. Then

$$(A_0 - B_0D_0^{-1}C_0)^{-1} = A_0^{-1} + A_0^{-1}B_0(D_0 - C_0A_0^{-1}B_0)^{-1}C_0A_0^{-1}.$$

Now we apply this lemma with the following choices:

$$A_0 = I, \quad B_0 = C, \quad C_0 = -K, \quad D_0 = qI - A.$$

Then we get

$$(I + C(qI - A)^{-1}K)^{-1} = I - C(qI - A + KC)^{-1}K.$$

□

Chapter 9

Identification of AR processes

9.1 Least Squares estimate of an AR process

Exercise 9.1 *Prove that R^* is non-singular.*

Solution: Obviously R^* is positive semidefinite, as it is the covariance matrix of $(y_{n-1}, \dots, y_{n-p})$. The singularity of R^* would imply the singularity of the process y . \square

Exercise 9.2 *Prove that R^* is non-singular by taking a state-space representation of y .*

Hint: Show that the pair (\tilde{A}, b) , as defined in Section 7, is controllable.

Exercise 9.3 * *Show that Condition*

$$\lim_N (-r_N) = \lim_N \frac{1}{N} \sum_{n=1}^N \varphi_n y_n = E\varphi_n y_n = -r^* \quad \text{a.s.}$$

implies that

$$\lim_N \frac{1}{N} \sum_{n=1}^N \varphi_n e_n = E\varphi_n e_n = 0 \quad \text{a.s.} \quad (9.1)$$

9.2 The asymptotic covariance matrix of the LSQ estimate

Exercise 9.4 *Assume that (e_n) is Gaussian. Show that $(\sum_{n=1}^N y_{n-1} y_{n-1})^{-1}$ has no finite expectation.*

Solution: We have

$$\sum_{n=1}^N y_{n-1} y_{n-1} \leq C \sum_{n=1}^N y'_{n-1} y'_{n-1},$$

where y'_n is an i.i.d. sequence of standard normally distributed random variables, hence the latter sum has a χ^2 distribution with N degrees of freedom. \square

9.3 The recursive LSQ method

Exercise 9.5 Assume that A and D are non-singular. Then

$$\left. \frac{d}{d\varepsilon} (A - \varepsilon BD^{-1}C)^{-1} \right|_{\varepsilon=0} = A^{-1}BD^{-1}CA^{-1}.$$

Solution: If $A = A(\varepsilon)$ is nonsingular, differentiable function of ε , then

$$\frac{d}{d\varepsilon} (A(\varepsilon))^{-1} = (A(\varepsilon))^{-1} \frac{d}{d\varepsilon} A(\varepsilon) (A(\varepsilon))^{-1}.$$

Now $A(\varepsilon) = A - \varepsilon BD^{-1}C$, thus

$$\frac{d}{d\varepsilon} (A - \varepsilon BD^{-1}C)^{-1} = (A - \varepsilon BD^{-1}C)^{-1} (A - \varepsilon BD^{-1}C) (A - \varepsilon BD^{-1}C)^{-1}.$$

Substituting $\varepsilon = 0$ we get the proposition of the exercise. \square

Exercise 9.6 Prove the Sherman-Morrison lemma.

Solution: This is a special case of the matrix inversion lemma. We consider the following block-matrix:

$$F = \begin{pmatrix} A & b \\ c^T & -1 \end{pmatrix}.$$

Then we have

$$\begin{aligned} (A - b(-1)c^T)^{-1} &= A^{-1} + A^{-1}b((-1) - c^T A^{-1}b)^{-1} c^T A^{-1} = \\ &= A^{-1} + A^{-1}bc^T A^{-1} \frac{-1}{1 + c^T A^{-1}b}. \end{aligned}$$

\square

Chapter 10

Identification of MA and ARMA models

10.1 Identification of MA models

Exercise 10.1 Prove that the stability of $C(z^{-1})$, implying the stability of \tilde{C} , yields that

$$\mathbb{E} |x_n^*(\theta) - x_n(\theta)|^2 = O(\gamma^n)$$

with any γ such that $\gamma > \varrho(\tilde{C})$, with $\varrho(\tilde{C})$ denoting the spectral radius of \tilde{C} (known to be less than 1).

Solution: Use the state-space representations for both processes, and take the difference of the two state-space equations. Exploit linearity to see the effect of the differences in initial values.

Exercise 10.2 Provide an expression of the coefficients c_l in terms of the roots, say γ_l , and express $\frac{\partial}{\partial \gamma}$ via $\frac{\partial}{\partial \theta}$.

Solution: The connection can be expressed via the following equation:

$$C(z^{-1}) = 1 + \sum_{l=1}^r c_l z^{-l} = \prod_{l=1}^r (1 - z^{-1} \gamma_l).$$

The derivative w.r.t. γ_l is:

$$\frac{\partial}{\partial \gamma_l} C(z^{-1}) = z^{-1} \prod_{j \neq l} (1 - z^{-1} \gamma_j) = \frac{z^{-1}}{1 - z^{-1} \gamma_l} C(z^{-1}).$$

10.2 The asymptotic covariance matrix of $\hat{\theta}_N$

Exercise 10.3 Show that the first term on the r.h.s. of the equality

$$\lim_N \frac{1}{N} V_{\theta\theta N}(\theta^*) = E \left(\varepsilon_{\theta\theta n}^*(\theta^*) e_n + \varepsilon_{\theta n}^*(\theta^*) \varepsilon_{\theta n}^{*T}(\theta^*) \right).$$

is zero.

Solution: As we have seen before $\varepsilon_{\theta\theta n}^*(\theta)$ is in the subspace spanned by $\{\varepsilon_{n-k}(\theta) : k = 1, 2, \dots\}$. For $\theta = \theta^*$ this set is H_{n-1}^e , and $e_n \perp H_{n-1}^y$. \square

10.3 Identification of ARMA models

Exercise 10.4 Show that if $A^*(z^{-1})$ and $C^*(z^{-1})$ have no common factor then R^* is non-singular.

Solution: Write $R^* = E \varepsilon_{\theta n}^{*T}(\theta^*) \varepsilon_{\theta n}^*(\theta^*)$, and assume that R^* is singular. Then there exists a pair of vectors, say (u^T, v^T) such that

$$(u^T, v^T) \varepsilon_{\theta n}^{*T}(\theta^*) = 0$$

a.s. Defining the polynomials $U(q^{-1})$ and $V(q^{-1})$ with their coefficients being equal to the components of u and v , resp., we arrive to the equality

$$\frac{U}{A^*} e - \frac{V}{C^*} e = 0$$

a.s. Rearranging this we get

$$UC^* - VA^* = 0,$$

which contradicts to the assumption that $A^*(z^{-1})$ and $C^*(z^{-1})$ have no common factor.

Exercise 10.5 Show that if $A^*(z^{-1})$ and $C^*(z^{-1})$ have a common factor then R^* is singular.

Solution: Apply the previous arguments in reverse order.

Exercise 10.6 *Compute the gradient process for the following models: MA(1), AR(1), ARMA(1,1).*

Solution:

MA(1) : The process is

$$y_n = e_n + c^* e_{n-1}.$$

The inverse process for $|c| < 1$ is:

$$\varepsilon_n + c\varepsilon_{n-1} = y_n.$$

The gradient process can be written as

$$\varepsilon_{c,n} + c\varepsilon_{c,n-1} + \varepsilon_{n-1} = 0, \quad \implies \quad \varepsilon_{c,n} + c\varepsilon_{c,n-1} = -\varepsilon_{n-1}.$$

AR(1) : The process is

$$y_n + a^* y_{n-1} = e_n.$$

The inverse process for $|a| < 1$ is:

$$\varepsilon_n = y_n + ay_{n-1}.$$

The gradient process can be written as

$$\varepsilon_{a,n} = y_{n-1}.$$

ARMA(1,1) : The process is

$$y_n + a^* y_{n-1} = e_n + c^* e_{n-1}.$$

The inverse process for $|a| < 1$ and $|c| < 1$ is:

$$\varepsilon_n + c\varepsilon_{n-1} = y_n + ay_{n-1}.$$

The gradient process can be written as

$$\varepsilon_{c,n} + c\varepsilon_{c,n-1} = -\varepsilon_{n-1},$$

$$\varepsilon_{a,n} + c\varepsilon_{a,n-1} = y_{n-1}.$$

□

Exercise 10.7 Show that for $\theta = \theta^*$ we have

$$\varepsilon_\theta(\theta)|_{\theta=\theta^*} = \left(\frac{1}{A^*} [q^{-1} \dots q^{-p}] e, -\frac{1}{C^*} [q^{-1} \dots q^{-r}] e \right)^T.$$

Solution: In **Lemma 10.10.** it was proved, that

$$C(q^{-1}) \varepsilon_\theta(\theta) = -\phi(\theta).$$

Substituting $\theta = \theta^*$ we get

$$C^*(q^{-1}) \varepsilon_\theta(\theta^*) = -\phi(\theta^*) = -(y_{n-1}, \dots, y_{n-p}, e_{n-1}, \dots, e_{n-r})^T. \quad (10.1)$$

Here we used $\varepsilon_n(\theta^*) = e_n$. The first part of the vector can be written as

$$(y_{n-1}, \dots, y_{n-p})^T = ([q^{-1} \dots q^{-p}] y)^T = ([q^{-1} \dots q^{-p}] \frac{C^*}{A^*} e)^T.$$

The second part of the vector is simply:

$$-(e_{n-1}, \dots, e_{n-r})^T = -([q^{-1} \dots q^{-r}] e)^T$$

Then dividing equation (10.1) by $C^*(q^{-1})$ we get the statement of the exercise. \square

Chapter 12

Stochastic volatility: ARCH and GARCH models

12.2 Stochastic volatility models

Exercise 12.1 *Show that the triplet $(\varepsilon_n, y_n, \sigma_n)$ itself is jointly strictly stationary.*

Solution: First we show that (y_n, σ_n) is jointly strictly stationary. Once we know $y_n, y_{n-1}, y_{n-2}, \dots$, any σ_k is a deterministic function of them for $k \leq n+1$. Moreover, the equation $\varepsilon_n = y_n/\sigma_n$ allows a deterministic computation of the error process if we know the other two. The process (y_n) is strictly stationary, thus the joint distribution of the history $y_n, y_{n-1}, y_{n-2}, \dots$ does not change under time shift. $(\varepsilon_n, y_n, \sigma_n)$ is obtained by a deterministic function from (y_n) , consequently their joint distributions are also insensitive to time shifts. \square

Exercise 12.2 *Show that under the conditions above*

$$E[y_n | \mathcal{F}_{n-1}^y] = 0 \quad \text{a.s.}$$

In other words (y_n) is a martingale difference process.

Solution: Write

$$E[y_n | \mathcal{F}_{n-1}^y] = E[\sigma_n \varepsilon_n | \mathcal{F}_{n-1}^y] = \sigma_n E[\varepsilon_n | \mathcal{F}_{n-1}^y] = 0.$$

The first equality is just the definition of y_n , the second uses the fact that σ_n is a function of the history of (y_n) . The is the result of (ε_n) being i.i.d. \square

Exercise 12.3 Show that under the conditions above

$$E[y_n^2 | \mathcal{F}_{n-1}^y] = \sigma_n^2 \quad \text{a.s.}$$

Solution: Following the lines of the previous solution we get

$$E[y_n^2 | \mathcal{F}_{n-1}^y] = E[\sigma_n^2 \varepsilon_n^2 | \mathcal{F}_{n-1}^y] = \sigma_n^2 E[\varepsilon_n^2 | \mathcal{F}_{n-1}^y] = \sigma_n^2.$$

□

Exercise 12.4 Show that under the conditions above (y_n) is a w.s.st. orthogonal process.

Solution: For $n > m$ we have

$$E[\sigma_n \varepsilon_n \sigma_m \varepsilon_m | \mathcal{F}_{n-1}^y] = E[\sigma_n \varepsilon_n \sigma_m \varepsilon_m | \mathcal{F}_{n-1}^\varepsilon] = \sigma_n \sigma_m \varepsilon_m E[\varepsilon_n | \mathcal{F}_{n-1}^\varepsilon] = 0.$$

□

12.4 State space representation

Exercise 12.5 Show that X_n defined by the right hand side of

$$X_n^* = u_n + \sum_{k=1}^{\infty} A_n A_{n-1} \dots A_{n-k+1} u_{n-k}$$

satisfies the state equation

$$X_{n+1}^* = A_{n+1}^* X_n^* + u_{n+1}, \quad n \in \mathbb{Z}$$

corresponding to the GARCH process.

Solution:

$$\begin{aligned} A_{n+1} X_n^* &= A_{n+1} u_n + A_{n+1} \sum_{k=1}^{\infty} A_n A_{n-1} \dots A_{n-k+1} u_{n-k} \\ &= \sum_{k=0}^{\infty} A_{n+1} A_n \dots A_{n-k+1} u_{n-k} = X_{n+1}^* - u_{n+1} \end{aligned}$$

□

Exercise 12.6 Prove uniqueness as stated in the theorem.

Hint: Let (X_{1n}) and (X_{2n}) be two solutions. Consider the dynamics for the difference process $(X_{1n} - X_{2n})$.

12.5 Existence of a strictly stationary solution

Exercise 12.7 Show that (u_n) is sub-additive, i.e. we have for any $n, m > 0$ the inequality $u_{n+m} \leq u_n + u_m$.

Solution: We have

$$u_{n+m} = E \log \|A_{n+m} \dots A_{n+1} A_n \dots A_1\|.$$

The norm is sub/multiplicative, i.e.

$$\|A \cdot B\| \leq \|A\| \cdot \|B\|.$$

Thus

$$\begin{aligned} \log \|A_{n+m} \dots A_{n+1} A_n \dots A_1\| &\leq \log(\|A_{n+m} \dots A_{n+1}\| \cdot \|A_n \dots A_1\|) \\ &= \log \|A_{n+m} \dots A_{n+1}\| + \log \|A_n \dots A_1\|. \end{aligned}$$

Taking expectation we get the statement of the Exercise. □

Lemma 4 Let (u_n) be a sub-additive sequence. Then the limit

$$\lambda = \lim_{n \rightarrow \infty} \frac{1}{n} u_n$$

exists, where $-\infty \leq \lambda < +\infty$, moreover

$$\lambda = \inf_{n \rightarrow \infty} \frac{1}{n} u_n.$$

Exercise 12.8 Prove the above lemma.

Solution: Let $\lambda = \inf_{n \rightarrow \infty} \frac{1}{n} u_n$. Take $\varepsilon > 0$ and choose a k such that $u_k/k < \lambda + \varepsilon$. Then for any integer n we have

$$u_{nk} \leq n u_k, \quad \implies \quad \frac{u_{nk}}{nk} < \lambda + \varepsilon.$$

Moreover, for any $0 \leq r < k$ and n large enough we get

$$\frac{1}{nk+r} u_{nk+r} \leq \frac{1}{nk+r} (u_{nk} + u_r) \leq \frac{1}{nk+r} (u_{nk} + \max_{0 \leq s < k} u_s) < \lambda + 2\varepsilon.$$

Consequently $\limsup_{n \rightarrow \infty} \frac{1}{n} u_n < \lambda + 2\varepsilon$. □

Exercise 12.9 Prove that for any $\varepsilon > 0$ there exist finite r.v.-s $C_n(\omega, \varepsilon)$ such that for any n we have

$$\|A_n A_{n-1} \dots A_{n-k+1}\| \leq C_n(\omega, \varepsilon) e^{(\lambda(\mathcal{A}) + \varepsilon)(n-k)}.$$

Show that $C_n(\omega, \varepsilon)$ can be assumed to be a stationary sequence.

Solution: After taking the logarithm we get the following constraint for $C_n(\omega, \varepsilon)$:

$$\log \|A_n A_{n-1} \dots A_{n-k+1}\| - (\lambda(\mathcal{A}) + \varepsilon)(n - k) \leq \log C_n(\omega, \varepsilon).$$

By the definition of the top-Lyapunov exponent the left hand side tends to $-\infty$ as $k \rightarrow \infty$.

Therefore

$$\sup_k (\log \|A_n A_{n-1} \dots A_{n-k+1}\| - (\lambda(\mathcal{A}) + \varepsilon)(n - k)) < \infty.$$

Choose this value as $C_n(\omega, \varepsilon)$. □

Exercise 12.10 Prove that for any sequence of r.v.-s (u_n) the condition

$$\sup_n \mathbb{E} \log^+ |u_n| < +\infty$$

implies that for any $\varepsilon > 0$ there exists a r.v. $C(\omega, \varepsilon)$ such that

$$|u_n| \leq C(\omega, \varepsilon) e^{\varepsilon n}.$$

Hint: Show that for any fix $C > 0$ the probability that the events $\{|u_n| > C e^{\varepsilon n}\}$ occur infinitely often is 0. For this use the Borel-Cantelli lemma.

Chapter 13

High-frequency data. Poisson processes

13.3 Poisson point processes on a general state space

Exercise 13.1 Show that μ is non-atomic, i.e. for all $x \in S$ we have $\mu(\{x\}) = 0$.

Solution: Clearly $N(\{x\}) = |\Pi \cap \{x\}|$ can only be 0 or 1. On the other hand, if we had $\mu(\{x\}) = \varepsilon > 0$, then $N(\{x\})$ would follow a Poisson distribution with a positive parameter. But this can take values larger than one with positive probability, which is a contradiction. \square

Exercise 13.2 Show that for any $A_1, A_2 \in \mathcal{G}$ with $\mu(A_i) < \infty$ we have

$$\text{cov}(N(A_1), N(A_2)) = \mu(A_1 \cap A_2).$$

Solution: Let $B = (A_1 \cap A_2)$. We may write the covariance as follows:

$$\text{cov}(N(A_1), N(A_2)) = \text{cov}(N(A_1 \setminus B) + N(B), N(A_2 \setminus B) + N(B)).$$

Using the bi-linearity of the covariance we can expand the r.h.s., and observe that all but one term vanishes due to independence. Thus

$$\text{cov}(N(A_1), N(A_2)) = \text{cov}(N(B), N(B)).$$

But the variable $N(B)$ follows a Poisson distribution with parameter $\mu(B)$, consequently its variance is indeed $\mu(B)$. \square

Theorem 2 *Let Π be a Poisson process with mean measure μ on S and let $S_1 \subset S$ be measurable. Then*

$$\Pi_1 = \Pi \cap S_1$$

is a Poisson process on S with mean measure

$$\mu_1(A) = \mu(A \cap S_1).$$

Exercise 13.3 *Prove the above theorem.*

Solution: We verify the requirements of the definition. For Π , (i) holds for any $A_1 \cap S_1, A_2 \cap S_1, \dots, A_n \cap S_1 \subset S$ provided that A_i are disjoint. But

$$N(A_i \cap S_1) = N_1(A_i),$$

thus (i) holds for Π_1 . Similarly

$$N_1(A) = N(A \cap S_1) \stackrel{\mathcal{L}}{=} P(\mu(A \cap S_1)) = P(\mu_1(A)).$$

□

Theorem 3 *Let Π be a Poisson process in \mathbb{R}^D with rate function $\lambda(x_1, \dots, x_D)$. Let $d < D$ and let*

$$\Pi_d$$

be the projection of Π on the first d coordinates. Then Π_d is a Poisson-process with rate function

$$\lambda^*(x_1, \dots, x_d) = \int \dots \int \lambda(x_1, \dots, x_D) dx_{d+1} \dots dx_D.$$

Exercise 13.4 *Prove the above projection theorem.*

Solution: Apply the mapping theorem with

$$f(x_1, \dots, x_D) = (x_1, \dots, x_d).$$

□

13.4 Construction of Poisson processes

Proposition 8 *Let X_1, \dots, X_n be i.i.d. on S according to the probability measure η . Assume that η has no atom. Then*

$$\Pi = \{X_1, \dots, X_n\}$$

is a Bernoulli process.

Exercise 13.5 *Prove the above proposition.*

Solution: Verify

$$P(N(A_0) = n_0, \dots, N(A_k) = n_k) = \frac{n!}{n_0! \dots n_k!} \eta(A_0)^{n_0} \dots \eta(A_k)^{n_k}$$

for the counts

$$N(A) = \#\{r : X_r \in A\}.$$

Use the fact that all X_i are different almost surely.

Exercise 13.6 *Work out the details of the proof of the existence of a Poisson process.*

Solution: First assume that $\mu(X) = \lambda < \infty$. Define $\eta(A) = \mu(A)/\mu(X)$. Choose the number of points in X , say N according to $P(\lambda)$. With N being fixed, construct the realization of a Bernoulli process. Then the resulting process will be a Poisson point process with intensity measure μ . The case $\mu(X) = \infty$ is handled by first writing

$$X = \cup_{i=1}^{\infty} X_i$$

so that $\mu(X_i) = \lambda_i < \infty$, (assuming that μ is σ -finite.)

□

13.5 Sums and integrals over Poisson processes

Exercise 13.7 *Prove the two identities:*

$$E(\Sigma) = \int_S f(x) \mu(dx).$$

and, if

$$\Sigma_1 = \sum_{x \in \Pi} f_1(x), \quad \Sigma_2 = \sum_{x \in \Pi} f_2(x),$$

then

$$\text{Cov}(\Sigma_1, \Sigma_2) = \int_S f_1(x) f_2(x) \mu(dx).$$

Solution: Both identities are true for step-functions. From there the final result follows by a simple limiting argument. \square

Exercise 13.8 Provide a formal proof for the above expression of $\text{var}(\Sigma)$ by differentiating the master equation

$$\mathbb{E}(e^{it\Sigma}) = \exp\left\{\int_S (e^{itf(x)} - 1) \mu(dx)\right\}$$

w.r.t. t once and twice, and setting $t = 0$.

Solution: Formal differentiation of the l.h.s. gives

$$\frac{d}{dt} \mathbb{E}(e^{it\Sigma}) = \mathbb{E}i\Sigma(e^{it\Sigma}) = \mathbb{E}i\Sigma,$$

when setting $t = 0$. Similarly, formal differentiation of the r.h.s. gives

$$\exp\left\{\int_S (e^{itf(x)} - 1) \mu(dx)\right\} \cdot \int_S if(x) e^{itf(x)} \mu(dx) = \int_S if(x) \mu(dx)$$

when setting $t = 0$, which proves the first claim. We proceed with the second derivatives similarly. \square

Exercise 13.9 Provide a formal proof for the above expression of $\text{cov}(\Sigma_1, \Sigma_2)$ by considering the master equation

$$\mathbb{E}(e^{it_1\Sigma_1 + it_2\Sigma_2}) = \exp\left\{\int_S (e^{it_1f_1(x) + it_2f_2(x)} - 1) \mu(dx)\right\}.$$

and taking mixed second order partial derivatives

$$\frac{\partial^2}{\partial t_1 \partial t_2},$$

and setting $t_1 = t_2 = 0$.

Solution: Follow the proof of the previous exercise. \square

Chapter 14

High-frequency data. Lévy Processes

14.1 Motivation and basic properties

Exercise 14.1 *Show that the characteristic function of L_t can be written in the form*

$$\mathbb{E}e^{iuL_t} = e^{t\psi(u)}.$$

Solution: Write $\mathbb{E}e^{iuL_t} = e^{\rho_t(u)}$. Since the increments are i.i.d. we have

$$\rho_t(u) + \rho_{\Delta t}(u) = \rho_{t+\Delta t}(u).$$

Since $\rho_t(u)$ is continuous in t , it follows that $\rho_t(u) = t\psi(u)$ with some $\psi(u)$, as stated.

Theorem 4 *Let (L_t) be a pure-jump Lévy process (having no Brownian motion component), with finite variation trajectories, defined by*

$$L_t = \int_0^t \int_{\mathbf{R}^1} x N(ds, dx). \quad (14.1)$$

Then

$$\mathbb{E} \left[e^{iuL_t} \right] = \exp \left[t \left(ibu + \int_{\mathbf{R}^1} (e^{iux} - 1) \nu(dx) \right) \right].$$

Exercise 14.2 *Provide a derivation of the above simplified version of the Lévy-Khintchine formula for processes with finite variation trajectories, using Campbell's theorem.*

Solution: For pure-jump Lévy processes with finite variation trajectories given by (14.1) the integrability condition

$$\int_{\mathbf{R}^1} \min(|x|, 1) \nu(dx) < \infty,$$

implies the integrability condition present in Campbell's theorem with $f(t', x) = x$ for $0 \leq t' \leq t$. Then we will have $\Sigma_f = L_t$. The expression given for $E[e^{iuL_t}]$ given by Campbell's theorem is then just the simplified form of the Lévy-Khintchine formula:

□